Do Index Funds Monitor?

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We thank Russell for providing index data. We thank conference and seminar participants at the ECGI, IESE Business School, University of Bolzano, University of Bristol, University of Colorado, University of Gothenburg, Lund University, University of Miami, Pennsylvania State University, Stockholm Business School, Tel-Aviv University, University of Toronto, University of Utah, WU Vienna, and York University, the 2019 American Finance Association annual meeting, the 2019 Financial Accounting and Reporting Section, and the 2019 Financial Intermediation Research Society conference. We also thank Yakov Amihud, Lucian Bebchuk, Tara Bhandari, Alon Brav, Alex Edmans, Amy Edwards, Yonca Ertimur, Fabrizio Ferri, Todd Gormley, Jillian Grennan, Peter Iliev, Wei Jiang, Avner Kalay, Zach Kaplan, Phil Mackintosh, Will Mullins, Gaizka Ormazabal, Stefan Petry, Martin Schmalz, Miriam Schwartz-Ziv, Nathan Seegert, Steve Stubben, Wei Wei, Toni Whited, and Alex Young for comments and suggestions. We thank Peter Iliev for sharing code and data with us.

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Abstract

Passively managed index funds now hold over 25% of U.S. mutual fund and ETF assets. The rise of index investing raises fundamental questions about monitoring and corporate governance. We show that, relative to active funds, index funds are less likely to vote against firm management on contentious corporate governance issues. Across a variety of tests, we also find no evidence that index funds engage with firm management to improve governance. Moreover, higher index fund ownership leads to less independent boards and worse corporate governance. Our results show that the rise of index investing is shifting control from investors to firm managers.

Keywords: Corporate Governance, Index Funds, Monitoring, Passive Investing, Voting

JEL Classifications: G12, G14

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I. Introduction

The separation of ownership and control generates agency conflicts between a firm’s managers and its shareholders. This well-known problem has been studied since at least the time of Adam Smith (1776).¹ Yet recently there has been a fundamental shift – index investing has changed the intermediary connecting investors to firms, potentially exacerbating this classic agency conflict. Over the last 25 years public corporations have experienced a dramatic increase in ownership by passively managed index funds (see Figure 1), and index funds are now the largest shareholders of many U.S. corporations (Azar, Tecu, and Schmalz (2018); Bebchuk and Hirst (2019b)). Although the increasingly large positions held by index funds should motivate them to monitor their portfolio firms (Grossman and Hart (1980), Shleifer and Vishny (1986), Admati, Pfleiderer, and Zechner (1994)), these new intermediaries have different incentives than traditional active mutual funds (Bebchuk, Cohen, and Hirst (2017), Edmans, Levit, and Reilly (2018)). As a consequence, the rise of index investing raises fundamental questions about monitoring and corporate governance. Specifically, to what extent do index funds monitor their portfolio companies?

In this paper, we study the monitoring behavior of index funds by examining the voice mechanism.² We find that index funds are 12.5 percentage points less likely to vote against firm management on contentious issues, compared to active funds. This result is particularly relevant given that index funds collectively vote over 25% of all shares in S&P 500 firms

¹Smith wrote, “The directors of such [joint stock] companies, however, being the managers rather of other people’s money than of their own, it cannot well be expected that they should watch over it with the same anxious vigilance with which the partners in a private copartnery frequently watch over their own.”

²Although economic theory indicates that there are two monitoring mechanisms, voice and exit, we note that the latter is less of an option for index funds since their objective is to reduce tracking error with the underlying index. However, since many index funds do not fully replicate their benchmark index they could, theoretically, exit a position. In the Appendix we examine the exit mechanism and show that index funds are significantly less likely to exit their portfolio firms, both unconditionally and after they lose a vote.
(Bebchuk and Hirst (2019b)). Furthermore, across a number of tests we find no evidence that index funds engage directly or indirectly with firm management. We also find that index funds do not improve corporate governance of their portfolio firms. Consistent with the theoretical predictions in Bebchuk et al. (2017) and Edmans et al. (2018), our results uniformly indicate that, relative to active funds, index funds are more likely to cede power to firm managers.

Given the increasingly large positions held by passively managed index funds, principal-agent theory would argue that these funds have strong incentives to monitor (Jensen and Meckling (1976), Grossman and Hart (1980), Shleifer and Vishny (1986), Admati et al. (1994)). Moreover, since the need to minimize tracking error makes it costly for index funds to exit a position, index funds should have strong incentives to enforce good governance through the voice mechanism (Bhide (1993), Fisch, Hamdani, and Davidoff Solomon (2018)). Consistent with this view, a number of recent studies argue that passively managed index funds are “closet activists” who improve a variety of corporate policies, from dividends and disclosure to competitive strategy.\(^3\)

However, the business model of passively managed index funds suggests that these funds have weaker incentives to monitor compared to traditional active funds, since they typically have a large number of firms in their portfolio and limited resources to invest in monitoring due to their low-cost structure. For example, in our sample the average index fund holds 371 stocks while the average active fund holds 116 stocks. Bebchuk and Hirst (2019a) point out that the top three index fund families have on average only 21 investment stewardship personnel to cover 17,849 firms in their portfolio. Moreover, index investing creates a free-
rider problem, because improvements to firm value are shared with all funds that follow the same index but the costs are borne only by the fund that exerts monitoring effort (Bebchuk et al. (2017)). To date, the net effect of index investing on firm governance is unclear. Our results provide novel evidence that index funds do not exert the same monitoring effort that active funds do.

The main challenge in studying fund monitoring behavior is that fund holdings are endogenous. First, firm characteristics such as size and liquidity jointly affect ownership and governance. Second, different firm policies attract different types of investors. Thus, there is the potential for endogeneity due to omitted variables and reverse causality. To address this concern we use a fixed effect structure that absorbs both unobserved firm heterogeneity and time-varying aggregate shocks. Moreover, because our voting analyses are conducted at the fund-firm-agenda item level, we can include firm-by-year fixed effects. Thus, we are even able to sweep out time-varying firm-level variation. Across all these specifications, our estimates remain highly stable, suggesting our results are not sensitive to influence by omitted confounders (Oster, 2019).

There is also the potential for selection bias: If a fund chooses not to hold a firm, we do not observe how that fund would have voted. To correct for selection bias, we develop a new research design that uses Russell index reconstitutions post-2006 to generate exogenous variation in fund holdings, and we use this to estimate a Heckman (1979) model. Our research design generates exogenous variation in index ownership in a difference-in-differences panel setting with firm fixed effects. Importantly, our difference-in-differences specification does not suffer from selection bias or bias due to noise in the forcing variable (Pei & Shen, 2017).

4Grinstein and Michaely (2005) find that higher firm payouts attract institutional holdings, while Brav, Jiang, Partnoy, and Thomas (2008) and Aghion, Van Reenen, and Zingales (2013) find that active investors target firms with weak governance and high leverage.
which are documented issues in prior studies that use Russell Index reconstitutions (see Wei and Young (2019)).

Our research design allows us to examine the effects of index investing on corporate governance in recent years, when index investing is most prevalent. However, there is another challenge to examining the monitoring behavior of index funds. A priori, it is unclear how index funds should behave if they are good monitors. Put differently, it is not clear what the null hypothesis is – for example, how often should index funds vote against managers? Our empirical design addresses this difficult challenge by comparing the monitoring behavior of index funds to the monitoring behavior of active funds. As a result, we provide the first definitive evidence that index funds monitor less than active funds.

We start by examining the voting behavior of funds across all agenda items. On consensus votes (i.e., when there is agreement between firm management and ISS—a third party proxy advisor) we find that index funds and active funds vote identically. By contrast, on contentious items (i.e., when ISS and firm management disagree), index funds are 12.5 percentage points more likely than active funds to vote with management. In a recent paper, Bebchuk et al. (2017) argue that index funds lack the incentives and resources to actively monitor their portfolio firms. Our findings provide novel evidence confirming this prediction. We find that index funds with low expense ratios are more likely to vote with management than index funds with high expense ratios. In other words, index funds with lower resources tend to invest less in costly monitoring (Lewellen and Lewellen (2018)); as a result, these

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5A recent paper by Wei and Young (2019) shows that the results in existing studies that use Russell Index reconstitutions are biased. In the attached Appendix on Russell Methodologies we confirm the problems in these existing studies and we provide extensive evidence that our results are not affected by these issues. Moreover, we are committed to sharing our code and data.

6In a further analysis, we examine voting at the fund family level, rather than at the individual fund level. Consistent with our main results, we find that fund families that have more assets under management (AUM) by index funds are more likely to vote with firm management.
funds are more likely to cede power to a firm’s management.

Arguably, not all votes are equally important. To shed further light on index funds’ monitoring behavior, we provide the first evidence on index fund voting across a wide variety of governance issues including: director elections, executive compensation, corporate disclosure, and managerial entrenchment. We find that index funds are more likely than active funds to vote with firm management on all of these agenda items. The results on managerial entrenchment are particularly relevant, since some of the largest index funds publicly state that they are against certain governance practices such as poison pills and golden parachutes. Yet, when it comes to voting on these issues, we find that index funds are actually less likely than active funds to vote against managerial entrenchment.

Overall, our voting results provide novel evidence that index funds cede power to firm management. Several recent studies reinforce our findings. For example, a contemporaneous working paper by Bolton, Li, Ravina, and Rosenthal (2019) finds that the “Big Three” fund families (Vanguard, BlackRock, and State Street Global Advisors, which hold a large portion of their assets in passive portfolios) rarely disagree with firm management, and Bebchuk and Hirst (2019a) provides descriptive statistics for voting on say-on-pay. They show that the Big Three fund families rarely vote “no” on executive compensation issues. Our analysis expands upon these in multiple dimensions: (1) we examine voting by all funds (instead of just the “Big Three”); (2) we examine voting on all governance issues (instead of examining one particular type of vote); (3) we develop an identification strategy to account for endogeneity concerns (while these papers only examine summary statistics); and (4) we use active funds as a counterfactual (which allows us to have a benchmark for how funds should vote if they are monitoring). Overall, across a wide variety of voting analyses, we consistently find novel

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evidence that index funds monitor less than active funds.

However, while our voting results suggest that index funds are weak monitors, it is possible that index funds monitor through other channels. For example, index funds could engage with managers either publicly or privately, and then vote in support of management proposals that they negotiated beforehand.\textsuperscript{8} We investigate index funds’ engagement in three ways. First, we split agenda items into shareholder proposals and management proposals (Gillan and Starks (2000)). Behind the scenes engagement could explain index funds’ voting behavior on management proposals, since they could be supporting management proposals that they negotiated beforehand. Yet the same reasoning does not apply on shareholder proposals. In other words, if index funds affect firm governance through private engagement, their tendency to vote with management should be mostly or entirely on management proposals and not on shareholder proposals. Yet we find that relative to active funds, index funds are significantly more likely to vote with firm management on both management and shareholder proposals.

Second, if index funds engage with the managers of their portfolio firms to implement governance changes, then we would expect to see a change in the number and/or type of agenda items proposed at the annual meeting after an exogenous increase in index funds’ holdings. Specifically, we would expect to see fewer contentious management proposals. Yet we do not see a change in the number of either management proposals or shareholder proposals, or in the fraction of contentious proposals. These results are inconsistent with the hypothesis that index funds privately engage with the management of their portfolio firms.

Third, we examine publicly observable evidence of mutual funds’ intent to engage. Shareholders are required to disclose a holding above 5\% of the firm’s market capitalization via

\textsuperscript{8}In their survey, McCahery, Sautner, and Starks (2016) find evidence of behind the scenes intervention by institutional investors. However, they do not distinguish between active funds and index funds.
Schedule 13D, which allows the fund to engage with the firm, or Schedule 13G, which does not. We find that fund families with more AUM in index funds are much less likely to file Schedule 13D. Moreover, for a subset of 13D filings that we can match to individual funds, we find there is not one single instance of an index fund filing a Schedule 13D. The results are consistent with Morley (2019) who argues that index funds could incur legal and compliance issues if they publicly engage with their portfolio companies. Also, Bebchuk and Hirst (2019a) find that the Big Three fund families filed only nine Schedule 13Ds between 2007 to 2018. However, Bebchuk and Hirst (2019a) do not specifically examine whether these filings were by index funds or active funds, nor do they examine funds outside these three fund families. We expand on their results by examining the universe of all funds and by directly comparing the behavior of index funds and active funds.

Finally, we examine whether exogenous changes in index fund holdings lead to changes at the firm level in standard governance measures, namely the E-Index (Bebchuk et al. (2008)), the fraction of independent directors on the board, poison pills, shareholder restrictions, or dual class share structures (Appel et al. (2016)). We find that exogenous increases in index fund holdings are followed by zero or negative effects on all five measures of corporate governance.

Taken together, these findings all point to the same conclusion: Index funds have limited resources and incentives to invest in costly monitoring of their portfolio firms. It follows that they monitor less than active funds do. Across a wide variety of tests and specifications, the data uniformly indicate that index funds cede power to firm management. Index funds are more likely to vote with firm management, and there is no evidence that they engage with firm management to improve governance. In a further test we also find that, unsurprisingly, index fund do not use the exit channel either to improve governance. As a result, corporate
governance does not improve, and may worsen, in index funds’ portfolio firms.

Our paper contributes to the literature on agency conflicts and monitoring incentives arising from dispersed ownership (e.g., Berle and Means (1932); Jensen and Meckling (1976); Demsetz (1983); Shleifer and Vishny (1986); Admati et al. (1994); Burkart, Gromb, and Panunzi (1997)). Given the dramatic increase in ownership by passively managed index funds, and since index funds are now the largest blockholders of many U.S. corporations (Azar et al. (2018)), studying their monitoring behavior is of fundamental importance (Edmans (2014), Bebchuk and Hirst (2019a)). Our results provide novel and definitive evidence that the rise of index investing is shifting power from investors to firm managers.

Our paper also extends the literature on proxy voting by institutional investors. Cvijanović, Dasgupta, and Zachariadis (2016) find that mutual fund families vote with firm management more when the fund family is also a manager of the company’s pension plan. More recent studies examine the reliance of mutual fund families on proxy advisor recommendations and conclude that such recommendations have an influence over shareholders’ voting (e.g., Iliev and Lowry (2014); Larcker, McCall, andOrmazabal (2015); Malenko and Shen (2016)), while others describe mutual fund families’ voting ideology and their association with funds’ characteristics (Brav, Jiang, and Li (2018); Bolton et al. (2019)). We differ from these studies in that we focus on the monitoring behavior of index funds relative to active funds. Our approach allows us to document for the first time whether index funds monitor more, less, or the same amount as other professional investors. We conduct a systematic study of the voting and engagement behavior of index funds across a wide variety of channels, and develop a research design that addresses the endogeneity biases in this setting.

Finally, our paper contributes to a growing literature on the effects of index investing on

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There is also a sizable literature on the monitoring behavior of active investors (see, e.g., DeMarzo and Urošević (2006); Gillan and Starks (2007); Back, Collin-Dufresne, Fos, Li, and Ljungqvist (2018)).
firm behavior. Several previous studies have found strong positive effects of index investing on corporate governance and payouts (Appel et al. (2016); Crane et al. (2016); Appel et al. (2018)). However, several studies argue there are methodological issues with these results (Wei & Young, 2019; Gloßner, 2018). Our research design avoids these issues and uses more recent data from 2007 to the present. Our findings provide definitive evidence that index investing has negative effects on corporate governance.

II. Data and Summary Statistics

To examine the monitoring behavior of index funds, we combine data from the Center for Research in Security Prices (CRSP), Compustat, Institutional Shareholder Services (ISS), and the Frank Russell Company (Russell), as discussed in detail below, for the years from 2004 to 2018.

A. Data

We use the ISS Fund Voting data to measure mutual funds voting behavior. Starting from 2003, ISS records the votes cast by individual mutual funds and exchange traded funds (ETFs) at shareholder meetings for the majority of publicly traded U.S. firms. We link the ISS data by fund-year to the CRSP mutual fund database, requiring that all sample funds be U.S. equity funds with at least $10 million in assets under management. In Table I, we report summary statistics for all investment funds in our sample from 2004 to 2018. Relative

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10 One potential challenge for studies of fund voting is that funds incorporated as a trust, such as SPY and QQQ, are not subject to NP-X reporting requirements. As such, their voting data is not publicly reported anywhere. None of the Russell 2000 index funds are incorporated as trusts, so our voting results for Russell funds are not affected by the omission of this data. We thank Tara Bhandari and Amy Edwards at the Securities and Exchange Commission for helpful conversations on this topic.
to active funds, index funds are less numerous (5,700 fund-years vs 25,800 fund-years) and on average more diversified (371 stocks vs 116 stocks) with more assets under management ($3.3B vs $2.2B) and lower expense ratios (0.47% vs 1.16%).

We use Russell Index membership lists provided directly by Russell, and we match this data to firm and stock-level data from CRSP and Compustat. We measure fund holdings by combining the CRSP mutual fund holdings database with the Thomson Reuters S12 database. We find that both databases omit some holdings of certain mutual funds in certain years, but the omissions are largely orthogonal across the two databases. Combining the two databases yields good coverage of funds in all sample years. Formally, we take the union of the two databases; if a fund-firm-year holding is in one databases but not the other, we include it; if it is in both databases, we take the larger of the two positions. In unreported analyses, we find that all our results are similar when we use only S12 or only CRSP holdings data.

Table II Panel A reports summary statistics for the firm-years in our overall matched sample, which consists of all Russell 3000 firms (essentially all U.S. public stocks excluding extremely small microcap stocks) from 2004 to 2018. Panel B reports summary statistics for the Russell cohort sample, which consists of firm-years for firms that were nearby the yearly upper and lower Russell bands (see Section III for more details). We see that the Russell cohort sample consists of a subset of mid-size firms that are otherwise similar to the population of all firms. The average Russell cohort firm has a market capitalization of 2.5 billion dollars, total ownership by mutual funds of 9.56% of the firm’s market capitalization, and an entrenchment (“E”)-index of 3.2. The average ownership by index funds is 3.86% of

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**Notes:**
11 We do not impose any filters on firm or stock characteristics, because our identification strategy requires all firms that are in the Russell 1000 or Russell 2000 in cohort year $t$ and year $t - 1$.
12 For example, S12 omits some data on the Vanguard Russell 2000 fund, which is well covered in CRSP. Conversely, prior to 2008 CRSP omits some data on the iShares Russell 2000 fund.
market capitalization (0.93% of which is by Russell 2000 index funds, and 0.09% of which is by Russell 1000 index funds), and the average ownership by active funds is 5.70% of market capitalization.

B. Summary Statistics

We begin our analysis by examining the cross-sectional variation of voting outcomes between active and index funds using univariate summary statistics. Consistent with the literature, we define a passively managed index fund as a fund with fund flag “D” in the CRSP Mutual Fund Database, and we classify all other mutual funds as active funds (all variables are defined in Appendix A).\textsuperscript{13} Row 1 of Table III shows the distribution of fund votes across the entire set of agenda items (i.e., the full sample). Unconditionally, index funds vote \textit{Yes} 90.4% of the time compared to 89.4% of the time for active funds.

Many agenda items are largely procedural, such as voting to adjourn the meeting. Hence, we use the voting recommendations of the major proxy advisor (ISS) to understand agenda items for which there is disagreement with firm management. Accordingly, in the next four rows of Table III we split agenda items into two categories: (i) consensus items – items for which firm management and ISS made the same recommendation (rows 2-3), and (ii) contentious items – items for which firm management and ISS made opposing recommendations (rows 4-5). There is disagreement in the literature on whether proxy advisory firms follow a \textit{one-size-fits-all} approach in issuing their recommendations. While findings in Iliev and Lowry (2014) and Larcker et al. (2015) seem to support this view, Ertimur, Ferri, and Oesch (2013) conclude that proxy advisors do not appear to follow a \textit{one-size-fits-all} approach.

\textsuperscript{13}In CRSP, a fund with flag D is a “pure index fund” whose “objective is to match the total investment performance of a publicly recognized securities market index.” In unreported tests, we classify funds according to their fund name or their active shares (Petajisto (2009)) and our results are similar.
We abstract away from this discussion and base our analyses on the observation that proxy advisors fulfill a key economic role in processing a substantial amount of information helping institutional investors to make informed voting decisions, while voting in agreement with management recommendation is the least costly way to satisfy the investors fiduciary duty to cast shareholder votes.\textsuperscript{14} Hence, we use ISS recommendations simply as a way to identify contentious issues with firm management, and then examine mutual funds voting behavior.

We find that for items that firm management and ISS both approve, index funds vote \textit{Yes} 95.6\% of the time while active funds vote \textit{Yes} 96.0\% of the time. Similarly, for items that firm management and ISS both oppose, index funds vote \textit{Yes} 4.2\% of the time while active funds vote \textit{Yes} 5.1\% of the time. The rates at which active and index funds vote no, abstain, or fail to record a vote are also similar. Thus, on consensus votes, index funds and active funds vote identically. On contentious items the results are very different. For items that firm management supports but ISS opposes, index funds vote \textit{Yes} 54.3\% of the time compared to 41.9\% for active funds. For items that firm management opposes but ISS supports, index funds vote \textit{No} 53.5\% of the time compared to 46.0\% for active funds. Thus, in both cases index funds are significantly more likely to side with firm management. Summing across all contentious items and coding abstentions as “no” votes (consistent with Del Guercio, Seery, and Woidtke (2008)), index funds voted with firm management 55.5\% of the time while active funds voted with firm management 46.2\% of the time.

Interestingly, index funds are \textit{less} likely than active funds to abstain on contentious items, especially items that were supported by firm management but opposed by ISS. If a fund wishes to maintain its relationship with firm management, voting “abstain” may be

\textsuperscript{14}For Example Cvijanović et al. (2016), and Davis and Kim (2007) argue that business ties (i.e., managing a company’s pension fund) influence mutual fund’s pro-management voting. See Bebchuk et al. (2017) for a detailed discussion on this point).
preferred to voting against management’s recommendation (Bebchuk et al. (2017)). Openly voting against firm management carries a higher cost because it may damage the relationship between the investor and firm management (Cvijanović et al. (2016)). Since most items require a majority of all votes cast to approve a measure, abstention can have the same effect as voting against a proposal but be perceived as a “soft no” (Del Guercio et al. (2008)). Hence, finding that active funds abstain more often than index funds on contentious items again suggests that active funds are more likely to oppose firm management than index funds are. Overall, these descriptive results suggest that on contentious issues, index funds consistently vote with management more often than active funds do. These results are particularly relevant since in a recent survey McCahery et al. (2016) document that voting against management is a primary channel through which mutual funds can use their voice and change corporate policies.

III. Research Design

A. Fixed Effects Structure

The voting statistics clearly show that on contentious items – agenda items on which firm management and ISS issued opposing recommendations – index funds were significantly more likely to side with firm management. However, a limitation of the results presented so far is that fund holdings are potentially endogenous with unobservable firm characteristics as well as time-varying aggregate shocks. For example, differences in firm characteristics (e.g., governance, entrenchment, managerial quality, etc.) may drive fund voting behavior. To address this concern, we use firm fixed effects which absorb time invariant differences across firms. We also include time fixed effects to absorb time-varying aggregate shocks.
However, it remains possible that corporate policies may change over time, which may change fund voting behavior. To address this concern, we exploit the richness of the voting data which contains multiple observations per firm-date. As a result, we can estimate specifications that use firm-by-year fixed effects (and in robustness checks, fixed effects by individual agenda item), which compare voting by active and index funds at the same annual meeting. Put differently, these specifications absorb even time-varying firm characteristics. As a result, they address any possibility of confounding variables that could bias our comparisons. We also stress that our analyses compare the behavior of index funds relative to active funds, which also helps mitigate the potential impact of confounding variables.

Of course, since funds choose which firms to hold, there is still the potential for selection bias. If a fund chooses not to hold a firm, then we do not observe how the fund would have voted. Thus, if index funds’ holdings differ systematically from active funds’ holdings that could bias any comparisons. To address this concern, we also implement a new research design that exploits post-2006 Russell index assignments in a Heckman (1979) correction model. In the next subsections we describe the post-2006 Russell assignment regime and the features of our research design.

B. Background on Russell Indexes

In June of each year Russell Investments reconstitutes their popular Russell 1000 (large-cap) and Russell 2000 (small-cap) indexes. To determine index assignment, Russell ranks all qualifying U.S. common stocks by their market capitalization as of the last business day in
May. Before June 2007, index assignment followed a simple threshold rule: Stocks ranked from 1-1000 were assigned to the Russell 1000 while stocks ranked from 1001-3000 were assigned to the Russell 2000.

Starting in June 2007, Russell implemented a new assignment regime ("banding"). After sorting stocks by their market capitalization, Russell computes an upper and lower band around the rank-1000 cutoff; the bands are calculated as +/- 2.5% of the total market capitalization of the Russell 3000E. Stocks within the bands do not switch indexes. That is, if a stock that was in the Russell 2000 last year is above the rank-1000 cutoff but below the upper band, it will stay in the Russell 2000 the following year, and vice versa.

Figure 2 plots index assignments in 2007, the first year of the banding regime. We observe that banding entirely eliminated the discontinuity across the rank-1000 cutoff; hence, an RDD across the cutoff is no longer feasible. However, Figure 2 also shows there are two new discontinuities at the upper and lower bands (dashed vertical lines). These discontinuities correspond to whether stocks switched indexes or stayed in their previous index. For example, consider a stock in the Russell 2000 that is nearby the upper band when the indexes are reconstituted. The stock’s index assignment depends on four parameters as calculated by Russell: 1) The stock’s ranking in the Russell 3000; 2) The market capitalization of the rank-1000 stock; 3) The total market capitalization of the Russell 3000E; 4) The cumulative market cap of the stocks ranked below the focal stock but above the rank-1000 stock. All four parameters are difficult to predict ex ante – indeed, Russell does not make their unadjusted

\[15\] Russell reports the index weights on the component stocks, which are based on their proprietary calculation of float-adjusted market capitalization. However, Russell does not disclose the initial rankings that determine index assignment, which are based on unadjusted market capitalization. We compute our own proxy market capitalizations and rankings at the end of May each year using CRSP and Compustat data following Chang, Hong, and Liskovich (2015). Our imputed Russell rankings recover the actual Russell index membership for 99.5% of firm-years, and all results are similar when we use alternative methods of imputing the Russell rankings.

\[16\] The 3000E is an “extended” version of the Russell 3000 that includes microcap stocks.
market capitalizations or rankings available *ex post*. All four parameters are difficult or impossible to manipulate. This line of reasoning suggests that within a sufficiently narrow window around each band in each year, whether a stock ranks above or below the band – and therefore switches or stays – is as good as randomly assigned.

### C. Heckman Correction

For each Russell index reconstitution since June 2007, we select a *cohort* that consists of two sets of treated and control stocks. Specifically, we select all stocks that were potential switchers (based on their lagged index membership) in windows of +/-100 ranks around the upper and lower band. Consider for example two stocks A and B that are similar in every way, including that both are in the Russell 1000 index in the year prior to treatment. Both stocks experience negative returns in the year prior to treatment and fall in the rankings. Firm A’s market capitalization falls by 10% while Firm B’s market capitalization falls by 10% plus epsilon. As a result, stock A stays in the Russell 1000 (and serves as a control), whereas stock B crosses the lower band and switches to the Russell 2000 (and is treated). Importantly, our identification strategy compares stocks that started in the same index and are similar in every dimension, including their lagged returns, except that they barely landed on different sides of the same band. Figure 3 shows the treated and control stocks around both bands in the 2007 cohort.

The fact that Russell index membership generates a discontinuity in treatment status suggests a regression discontinuity design (RDD). However, there are features of the setting that make an RDD undesirable. The main feature is that we do not observe the true rankings that determines index assignment; instead, we must impute them using the CRSP and Compustat data. This is a concern because errors in measuring the forcing variable
bias the RDD control function to be too flat, and produce spurious or biased estimates of treatment effects (Pei & Shen, 2017).\(^\text{17}\) To deal with this issue, we estimate a cohort difference-in-differences specification with firm-by-cohort fixed effects. For each stock in each cohort, we include firm-years from three years prior to the cohort year (pretreatment years -3, -2, -1) and three years after the cohort year (post-treatment years 0, 1, 2) in the sample. Formal balance tests show that the treated and control firms are indistinguishable \textit{ex ante} across both bands on every dimension.\(^\text{18}\) We measure the effects of switching indexes using the following difference-in-differences specification:

\[
Y_{jt} = \beta_1 I\{R1000 \rightarrow R2000_{jc}\} \times \text{PostAssignment}_{ct} \\
+ \beta_2 I\{R2000 \rightarrow R1000_{jc}\} \times \text{PostAssignment}_{ct} \\
+ \phi_{jc} + \lambda_t + \epsilon_{jct},
\]

where \(\phi_{jc}\) and \(\lambda_t\) are, respectively, firm-by-cohort and year fixed effects, \(\text{PostAssignment}\) is an indicator variable that equals 1 after index assignment, \(R1000 \rightarrow R2000\) equals 1 if a firm switched from the Russell 1000 to the Russell 2000, whereas \(R2000 \rightarrow R1000\) equals 1 if a firm switched from the Russell 2000 to the Russell 1000.\(^\text{19}\) We compare firm outcomes before treatment versus after treatment, with a fixed effect applied to each firm in each cohort. Because each firm had a single, fixed ranking within a cohort, the fixed effects \(\phi_{jc}\) absorb any correlation of the outcome variable with both the true ranking and the error in the proxy ranking for each firm. Thus, the specification (1) estimates the effects of

\(^{17}\)Note that a fuzzy RDD, which adjusts for non-compliance with treatment assignment, does not address this issue. We thank Toni Whited for pointing this out.

\(^{18}\)See the Appendix on Russell Methodologies for further details.

\(^{19}\)Importantly, this means that \(\beta_1\) and \(\beta_2\) – the effects of switching from the R1000 to the R2000 and vice versa – are identified from disjoint sets of treated and control stocks. The firm-by-cohort fixed effects sweep out any time invariant differences between treated and control stocks, while the year fixed effects remove aggregate trends in firm behavior or ownership.
switching indexes, as would a perfectly measured RDD, but in a way that is not sensitive to measurement error in the forcing variable. Our difference-in-differences approach with firm-by-cohort fixed effects addresses error in the forcing variable because it eliminates the need for a control function. Indeed, any control function would be subsumed by the firm-by-cohort fixed effects.

Our methodology differs from previous papers that use Russell index reconstitutions in three main ways. First, our paper is the first to focus exclusively on the post-2006 period and to develop a research design specifically around the post-banding assignment regime. Thus, our results reflect this more recent period, during which index investing is at all-time highs. Second, unlike previous RDD approaches, our difference-in-differences approach uses high dimensional fixed effects to sweep out observed and unobserved heterogeneity among firms. Among other advantages, this means that our estimates are not biased by noise in the measurement of the forcing variable, which can be an issue in RDD estimates (Pei & Shen, 2017). Third, we then use our cohort difference-in-differences specification as the first stage in a Heckman (1979) correction model, to address potential selection bias. Specifically, we estimate the following two-stage model:

\[
\text{Observed}_{ijt} = \text{Probit}(\tau \text{IndexFund}_i \\
+ \xi_1R1000 \rightarrow R2000_{jc} \times \text{PostAssignment}_{ct} \times \text{IndexFund}_i \\
+ \xi_2R2000 \rightarrow R1000_{jc} \times \text{PostAssignment}_{ct} \times \text{IndexFund}_i \\
+ \mu_1 R1000 \rightarrow R2000_{jc} \times \text{PostAssignment}_{ct} \\
+ \mu_2 R2000 \rightarrow R1000_{jc} \times \text{PostAssignment}_{ct} \\
+ \phi_{jc} + \chi_t + \nu_{ijct})
\]

\(2\)

E.g., Mullins (2014); Chang et al. (2015); Crane et al. (2016); Appel et al. (2016).
\[ Y_{ijt} = \beta \text{IndexFund}_i + \alpha \text{InverseMillsRatio}_{ijt} + \lambda_j + \kappa_t + \epsilon_{ijt} \] 

(3)

Equation (2) uses our cohort difference-in-differences specification to generate exogenous variation in fund ownership. \( \text{Observed} \) equals 1 if fund \( i \) holds stock \( j \) on date \( t \), and zero otherwise; \( \text{IndexFund} \) equals 1 if the fund is an index fund, and 0 otherwise; \( R1000 \rightarrow R2000 \) equals 1 if a firm switched from the Russell 1000 to the Russell 2000, whereas \( R2000 \rightarrow R1000 \) equals 1 if a firm switched from the Russell 2000 to the Russell 1000. \( \text{PostAssignment} \) equals 1 if the firm-year is post index assignment, and 0 if it is pre index assignment. \( \phi_{jc} \) and \( \chi_t \) are, respectively, firm-by-cohort and year fixed effects. The results for the first stage (Equation (2)) are reported in Appendix B, Table A1.

Equation (3) shows the second stage, which examines outcome variables as a function of index fund ownership, after including the \( \text{InverseMillsRatio} \) (i.e., the Heckman correction term from Equation (2)). \( \lambda_j \) are firm fixed effects and \( \kappa_t \) are year fixed effects.

D. Effects of Index Switching on Fund Ownership

We next examine whether Russell index assignment changes fund ownership. This is a necessary condition for our Heckman (1979) model. In Column 1 of Table IV, we present estimates of the effects of index switching on ownership by Russell 2000 index funds using our cohort difference-in-differences specification (equation (1)).\(^{21}\) We find that ownership by Russell 2000 index funds rises by an average 1.45\% of market capitalization for stocks that switch into the Russell 2000 relative to similar stocks just above the lower band that stay in the Russell 1000. At the same time, we find that ownership falls by 1.34\% of market capitalization.

\(^{21}\)The table separately examines fund holdings based on each fund’s benchmark index: Column (1) examines ownership by Russell 2000 funds only (31 funds), Column (2) examines ownership by Russell 1000 funds only (32 funds), and Column (3) examines ownership by all index funds combined.

Electronic copy available at: https://ssrn.com/abstract=3259433
capitalization for stocks that switch into the Russell 1000 relative to similar stocks just below the upper band that stay in the Russell 2000. The two coefficient estimates are similar in magnitude, even though they are estimated from two disjoint sets of stocks.

In Column 2 of Table IV, we report the effects of index switching on ownership by Russell 1000 funds. As expected, we find the opposite effect (relative to the change in ownership by Russell 2000 funds shown in Column 1). However, the coefficient is smaller for Russell 1000 fund holdings, falling by 0.18% of market capitalization in the lower band treatment group and rising by 0.17% of market capitalization in the upper band treatment group. This is as expected, because the index weights of stocks at the bottom of the Russell 1000 are significantly smaller than the index weights of stocks at the top of the Russell 2000. Similar to Column 1, the two coefficient estimates are similar in magnitude, even though they are estimated from two disjoint sets of stocks.

The net effect on holdings by all index funds in the data (Table IV Column 3) is similar to the net effect on holdings by Russell 1000 and 2000 index funds. By contrast, in Column 4 we examine the effects of index switching on ownership by active mutual funds. The changes in ownership by active funds are small and not statistically significant. As a result, total holdings by all mutual funds (Table IV Column 5) are entirely driven by changes in holdings by index funds.

Figure 4 plots Russell 2000 fund ownership for the four groups (switchers vs. stayers near the upper band; switchers vs. stayers near the lower band) in event time, that is, the observation year minus the cohort year. The results clearly show that: (i) Switchers and stayers in both groups have the same pre-treatment levels and trends, and (ii) switching into the Russell 2000 leads to higher index fund ownership and vice versa. Because firms in any group may also switch indexes in the post-treatment years, the treated and control
groups converge toward each other after the treatment year. In sum, the evidence shows that index switching is plausibly random among sample firms near the yearly Russell bands, and is followed by symmetric shifts in ownership by index funds.

IV. Voting

In this and the next section, we examine the monitoring behavior of index funds; we start with OLS regressions across all firms and then move to our Russell cohort setting using a Heckman (1979) model. Our findings are the same across all specifications and samples.

In Table V, Columns 1 through 4, we estimate the difference in fund voting, on all contentious items, across the universe of all funds and firms. The dependent variable $Voted\text{-WithMgmt}$ equals 1 if a fund votes with management’s recommendation, and 0 otherwise.\footnote{Following management’s recommendation is defined as voting Yes on a recommendation of Yes, and No or Abstain on a recommendation of No or Withhold.} The independent variable $IndexFund$ equals 1 if the fund is an index fund, and 0 if the fund is an active fund. The estimates in Columns 1 and 2 include firm fixed effects, which remove time invariant differences across firms, and year fixed effects which remove aggregate trends. In Column 1 we find that, compared to active funds, index funds are 12.5 percentage points more likely to side with firm management on contentious items. This is a larger difference than that in the summary statistics (Table III) and is due to the addition of firm fixed effects, so that we now compare voting by index funds to active funds within each firm.

Lewellen and Lewellen (2018) argue that there is a direct effect of a fund’s expense ratio on their incentive to monitor. Thus, in Column 2 we add as an explanatory variable each fund’s yearly expense ratio. We estimate the coefficient on the expense ratio separately for index and active funds because of the different incentives and expense ratios that the two
types of funds have. We find that active funds’ voting behavior does not vary significantly with their expense ratio. This could indicate that active fund expense ratios are always high enough to provide monitoring incentives. By contrast, index funds with higher expense ratios are less likely to side with firm management on contentious items. The coefficient of $-0.238$ means that an index fund with an expense ratio that is 25 basis points higher (about one standard deviation) is 6.1 percentage points less likely to side with firm management, which is half of the overall difference in voting between index and active funds. This result is strikingly consistent with the prediction of Bebchuk et al. (2017): The economics of index investing restrict the resources that a fund has to employ in monitoring, since index funds compete on providing a standardized product at the lowest price. In other words, when index funds have more resources to employ in monitoring, they behave more like active funds (i.e., they side less with firm management on contentious items).

The firm and year fixed effects in Columns 1 and 2 mitigate concerns of an omitted variable bias due to a time-invariant characteristic or an omitted aggregate shock. However, an additional concern is that corporate policies change over time, which may be related to fund voting behavior. Hence, in Columns 3 and 4, we compare fund voting using firm-by-year fixed effects. This approach sweeps out time-varying firm characteristics. We find that the difference in voting between index funds and active funds is nearly identical (12.4% vs 12.5%) as is the relation with fund expense ratios. These results suggest that firm and year fixed effects, as in Columns 1 and 2, account for most or all of the unobserved heterogeneity that could bias our treatment effect. In other words, voting-relevant characteristics such as firm governance, managerial quality, or shareholder engagement vary widely between firms.
but vary little within firms over time.\textsuperscript{23}

For a variety of estimates that cover the universe of funds and firms in our data, we find that index funds are more likely to vote with firm management on contentious items. However, there is still the potential for selection bias, because funds choose which firms they hold. If index funds tend to hold better-run firms, or vice versa, then the gap in fund voting behavior might be explained by their holdings and not by their monitoring. To explicitly correct for selection bias in fund holdings, we implement a Heckman (1979) approach in our Russell setting. First, for comparison, in Columns 5 and 6 of Table V we present results for the Russell subsample without the Heckman correction. We find similar results to those in the entire sample. Next, in Columns 7 and 8 of Table V we report the second-stage estimates according to the Heckman (1979) model described in equation (3).

The gap in voting behavior between index and active funds is still present: Index funds are 8.4 percentage points more likely than active funds to side with firm management, and again index funds with higher expense ratios are significantly less likely to side with firm management. These results suggest that part of the gap in voting behavior is due to selection: Active funds may choose to hold firms whose management they are more likely to disagree with relative to index funds. But, even after correcting for selection bias, we again find there is a statistically and economically significant difference in the voting behavior between active and index funds.\textsuperscript{24}

\textsuperscript{23}Appendix C presents alternate specifications including estimates with fixed effects for each individual agenda item, that is, comparing how funds voted on the same agenda item. The results are again similar to the specifications with firm and year fixed effects.

\textsuperscript{24}In many cases, funds belong to fund families such as Fidelity or Vanguard, and voting might be decided at least partly at the fund-family level or at the investment advisor level. For example, funds like Vanguard have a centralized voting authority, while others like BlackRock have a more decentralized approach. Such coordination is clear in the data: We find that the fund-family identity explains 26\% of the variation in fund voting, while fund identity (which is nested within fund-family identity) explains 33\%. Appendix C presents results when we examine voting policy at the fund-family level. Again, the conclusion is the same: Funds in families with more AUM in index funds are more likely to vote with firm management.
In sum, across a wide range of specifications and samples, we consistently find the same result: Index funds are more likely than active funds to vote with firm management on contentious agenda items. Moreover, index funds with higher expense ratios are less likely to vote with firm management – that is, they vote more like active funds do. The results suggest that index funds lack the incentives and resources to regularly monitor their portfolio firms, consistent with the predictions in Bebchuk et al. (2017). More generally, these findings shed light on a related question: If index funds exert little monitoring effort, why don’t they let ISS decide their votes 100% of the time? As noted in Del Guercio et al. (2008) and Bebchuk et al. (2017), it is costly for some class of shareholders to oppose firm management. If index funds always vote with ISS, on contentious issues they would vote against a firm’s management, which may have long-term implications for their relationship with the firm management team. While index funds have little incentive to disagree with firm management, it is also costly to damage their relationship with firms’ managers (Cvijanović et al. (2016); Davis and Kim (2007)). Taking all the evidence together, the implications are clear: (i) For index funds, the benefit of monitoring is low, (ii) the cost of monitoring is positive, (iii) disagreeing with a firm’s management adds additional costs. Accordingly, unless the benefit for voting against management is incredibly large, index funds choose to vote in accordance with managerial preferences. In other words, owing to their incentives, index funds cede power to firm management.

25 Consistent with this, in the prospectus for Vanguard Index Funds dated April 25, 2018 Vanguard states: “We will give substantial weight to the recommendations of the company’s board, absent guidelines or other specific facts that would support a vote against management.”
A. Voting on Agenda Items by Category

The results in Table V indicate that, on average, index funds are more likely than active funds to vote with firm management on contentious issues. Arguably, not all votes are equally important and sometimes votes may be procedural and nonbinding.\textsuperscript{26} Hence, a relevant subsequent question is whether the gap in voting behavior between active and index funds varies across different corporate governance policies. Accordingly, we examine how fund voting differs across the following categories of agenda items related to corporate governance where management and ISS disagree:

1. Board of Directors: Items whose description includes “director” or “board”; 
2. Compensation: Items whose description includes “compensation” or “incentive”. This category is mostly (83%) made up of say-on-pay votes; 
3. Disclosure: Items whose description includes “disclosure” or “reporting”; 
4. Entrenchment: Items whose description includes “staggered”, “bylaw”, “poison pill” or “parachute”.

We report results in Table VI. In Column 1, we find that index funds are 13.2 percentage points more likely to side with firm management on contentious items relating to the board of directors. A small subset of these items relate to formal proxy battles between the incumbent board and an activist shareholder. That is, our results in Column 1 are consistent with those of Brav et al. (2018), who focus on fund voting in proxy battles, and again consistent with the idea that index funds cede power to firm management on contentious governance issues.\textsuperscript{26}

\textsuperscript{26}However, prior studies have shown that a vote against management has a significant impact on firm policies and value, even when votes are nonbinding. For example, Ertimur et al. (2013) and Larcker et al. (2015) show that firms with substantial (actual or anticipated) voting dissent respond by removing controversial compensation practices.
Next, in Columns 2 to 4 we find that the gap in fund voting between index and active funds is positive and of similar magnitude for items related to compensation, disclosure, and managerial entrenchment. The results on managerial entrenchment are particularly relevant since some of the largest index funds publicly claim to be against certain governance practices such as poison pills and golden parachutes. Yet, voting in agreement with firm management on contentious items related to managerial entrenchment clearly indicates that index funds cede power to managers. More broadly, these findings suggest that the rise of index investing has consequences for (at least) board structure, managerial compensation, disclosure, and managerial entrenchment. In sum, relative to active funds, index funds cede power to firm management on decisions related to corporate governance across the board.

V. Engagement

An alternative hypothesis could allow for index funds to vote with firm management while still monitoring to ensure good corporate governance: Index funds could engage with the managers of their portfolio firms either publicly or privately (Dimson, Karakaş, and Li (2015); McCahery et al. (2016)), and then vote in support of management proposals that they negotiated beforehand. We test the engagement hypothesis in three ways. First, we split our estimates of fund voting behavior between agenda items that were proposed by management and those that were proposed by shareholders. Proposals by management could be affected by engagement, but proposals by shareholders should not. Thus, if index funds affect firm governance through private engagement, their tendency to vote with management should be mostly or entirely on management proposals and not on shareholder proposals. Second, if index funds engage with the managers of their portfolio firms to implement gov-

27See Bebchuk et al. (2008) and the ISG framework at https://isgframework.org.
ernance changes, then we would expect to see a change in the number and type of agenda items proposed at the annual meeting after an (exogenous) change in index fund holdings. Third, we study funds’ propensity to engage with the management of their portfolio firms by examining index funds’ propensity to file a Schedule 13D, which indicates that the fund intends to engage with the firm.

A. Voting on Management and Shareholder Proposals

A key test of the engagement hypothesis is whether the gap in voting between active funds and index funds depends on whether shareholders or management proposed the contentious agenda item. Proposals by management could be affected by engagement, but proposals by shareholders should not. Thus, if index funds affect firm governance through engagement, their tendency to vote with management should be mostly or entirely on management proposals and not on shareholder proposals. In Table VII, we split all contentious items in the sample accordingly. We find that the pattern that index funds are more likely to side with firm management holds true, regardless of who proposed the agenda item. On management proposals opposed by ISS, index funds are 14.4 percentage points more likely to vote “Yes.” Similarly, on shareholder proposals opposed by firm management, index funds are still 10.3 percentage points more likely to vote “No”, i.e. in support of firm management.

The results in Table VII address the concern that index funds might be voting in agreement with firm management after they engage with managers privately. Private engagement could apply to index funds’ voting on proposals by management, but it cannot apply to index funds’ voting on proposals by shareholders. Yet on contentious shareholder proposals, index funds again cede authority to firm management. These results are consistent with a contemporaneous paper by Brav et al. (2018). They show that in proxy contests, an impor-
tant category of shareholder proposals, index funds do not support activist shareholders but instead side with firm management.

Table VII presents another interesting observation: Index funds are significantly less likely than active funds to abstain on management proposals that were opposed by ISS. This finding is consistent with the argument of Del Guercio et al. (2008) and Bebchuk et al. (2017) that it is costly for shareholders to oppose firm management. It suggests again that index funds cede power to firm management, whereas active funds prefer to either directly oppose firm management (vote No) or to abstain (soft No).

B. Changes in Agenda Items at Shareholder Meetings

We further investigate funds’ propensity to privately engage with the management of their portfolio firms by examining the effect of changes in index fund holdings on the number and types of agenda items that appear at the firm’s annual meeting. For this analysis we use the firm-year panel and estimate the difference-in-differences specification as in equation (1). This specification allows us to estimate the changes in the number and type of agenda items, and the proportion of agenda items supported by management and/or ISS, that are caused by exogenous changes in index fund ownership of a firm. If index funds engage with the managers of their portfolio firms to implement governance changes, then we would expect to see a change in the number and/or type of agenda items proposed at the annual meeting after an exogenous increase in index funds’ holdings. Specifically, we would expect to see fewer contentious management proposals.

Table VIII presents the results. Neither set of treated firms (i.e. firms that switch Russell indexes in either direction) significantly changed the number of agenda items at their annual meetings in the post-treatment period (Column 1). There was also no change in the number
of proposals by shareholders (Column 2) or by firm management (Column 3). Moreover, we observe no change in the fraction of agenda items that were opposed by ISS (Column 4) or opposed by management (Column 5). Finally, if index funds privately engage with firm management, we would also expect to see an increase in the fraction of agenda items that were approved by both ISS and management. We see no such change (Column 6). It is important to note that the absence of any significant results is not due to power issues: The standard errors in each case are small enough to rule out any effect up to a fraction of one standard deviation.

In sum, our results are inconsistent with the private engagement hypothesis. Moreover, our findings complement the evidence in Bebchuk and Hirst (2019a) that index funds do not meet with the majority of their portfolio firms and accord with a recent working paper by Iliev, Kalodimos, and Lowry (2018) who document that, relative to active funds, index funds conduct significantly less research about their portfolio firms. Again, our results suggest that compared to active funds, index funds have weaker incentives to monitor and therefore cede power to the managers of their portfolio companies.

C. Disclosure: Schedule 13D vs Schedule 13G

Finally, we examine a public signal of the engagement hypothesis: The funds’ propensity to file a 13D schedule. The SEC requires shareholders to disclose a holding above 5% of any public company via either Schedule 13D or Schedule 13G. Schedule 13D is required if the shareholder has “the purpose or the effect” of influencing the control of the firm. This category includes actions such as proposing governance changes, or engaging with the portfolio company to propose or facilitate the appointment of particular individuals as directors, etc. The short-form Schedule 13G, by contrast, requires that the shareholder has no such

Table IX presents probit estimates of 13D filing behavior at the fund-family level. Because the Schedules 13D and G are filed at the level of the fund family, we match disclosure filings to fund families and not to the individual funds. In all, we match 30,864 disclosure filings since 2004 to a fund family in our data. The dependent variable is an indicator variable for whether each filing was under the “activist” Schedule 13D ($\text{Filed 13D}=1$) or the short-form and passive Schedule 13G ($\text{Filed 13D}=0$). The independent variable $\text{FracAUMPassive}_{jt}$ is the fraction of fund family $j$’s AUM that was managed by index funds in year $t$. Thus it ranges from zero for a fund family entirely populated by active funds, to one for a fund family entirely populated by index funds. Column 1 shows that fund families with more index fund AUM are significantly less likely to file Schedule 13D. The marginal effect (which corresponds to moving from 100% active to 100% index) is $-27$ percentage points, which is more than 100% of the base rate probability. The same conclusion holds when we control for the fund family’s total AUM (Column 2) and for the number of blockholding disclosures the family filed in that year (Column 3).

In a further step, we match blockholdings by individual funds, as revealed in the merged S12 and CRSP holdings data, to SEC disclosure filings by that fund’s parent family. We keep only matches that are unambiguous at the fund-firm-year level. In all, we match 4,475 disclosure filings to individual funds. For active mutual funds, 64 of 4085 filings were under Schedule 13D. For index funds, 0 of 390 filings were under Schedule 13D. In other words, for the sub-sample of filings that are verifiably by index funds, not a single fund filed form 13D.

Thus, fund families with more index fund AUM are less likely to file Schedule 13D and
more likely to file Schedule 13G. These findings are strongly supported by Morley (2019), who argues that index funds could incur legal, organizational and technical risks if they publicly engaged with their portfolio companies. Once again, our result are inconsistent with the hypothesis that index funds affect governance through engagement with their portfolio firms, but are instead consistent with the idea that, owing to their incentives, index funds cede power to firm management.

VI. Effects on Firm Governance

Our analysis so far has provided consistent evidence that index funds, compared to active funds, do not use their voice to change firm policies: they are less likely to vote against firm management on contentious governance issues, and there is no evidence that they either publicly or privately engage with the management of their portfolio firms. However, we note that index funds have repeatedly claimed to care about corporate governance. For example, the Big Three fund families often underline in their stewardship or proxy voting reports that their portfolio companies are aware of which governance principles they consider most important to the creation and protection of shareholder value over the long term. As a consequence, while we show that voting against management, filing a 13D or changing the number or types of agenda items at the annual meeting is not a priority for index funds, it could still be that managers of their portfolio companies change corporate governance in line with index funds claimed preferences as index fund ownership of their firms increases.

Accordingly, we examine the effects of index fund holdings on firm governance in our setting. The four governance policies generally highlighted by the Big Three fund families are: (1) to oppose entrenchment, which we measure by the E-index of Bebchuk et al. (2008); (2) to
support board independence, which we measure by the fraction of independent directors on
the board; (3) to oppose anti-takeover policies, which we measure by the removal of poison
pills and restrictions on the ability of shareholders to call a special meeting; and (4) to
oppose unequal voting rights, which we measure by dual-class share structure. We estimate
the difference-in-differences specification shown in equation (1), which allows us to estimate
changes in corporate governance policies that are caused by quasi-exogenous changes in fund
ownership. The results are shown in Table X.

In column 1, we find that when index fund ownership increases it is followed by an im-
provement in firm governance, as measured by a small decrease in the E-index. When index
fund ownership decreases, the effect on managerial entrenchment is positive but not signifi-
cant. In column 2, we find that when index fund ownership increases, board independence
decreases by 3% on average. In column 3 we find that when index fund ownership decreases,
poison pills are 4% more likely to be removed. Finally, in columns 4 and 5 we find no effect
of index fund ownership, in either direction, on shareholders’ ability to call a special meeting
or on the firms’ dual-class share structure.

These last four findings contrast with those of Appel et al. (2016), who find large positive
effects of index fund ownership on the same outcome variables. The Internet Appendix
Section IX provides a detailed comparison of both research designs. In particular, several
papers (e.g., (Wei & Young, 2019; Gloßner, 2018)) argue that the results in Appel et al. (2016)
are contaminated by selection bias, because they use the published (second-step) Russell
rankings instead of the unadjusted (first-step) rankings that determine index assignment.
Our results suggest that, instead, investment by passively managed index funds leads to
worse governance.

Overall, we examine the monitoring behavior of index funds as measured by voting and
engagement, and present a wide range of tests that all point to the same conclusion: Index funds have less incentive to monitor, so they monitor less than active funds do. Across a range of measures, we find that exogenous increases in index fund holdings are followed by small and generally negative effects on firm governance. Taken together, these results provide evidence that the voting behavior of index funds has consequences: Index funds do not improve, and may worsen, governance within their portfolio firms.

VII. Conclusion

Theoretically, the increasingly large positions held by index funds should motivate them to monitor their portfolio firms. Yet, these new intermediaries have different incentives than traditional active funds. To date, a large literature has examined the effect of index fund ownership on corporate outcomes without checking the direct relations between investors and firm management. We provide the first comprehensive evidence on this fundamental question. Specifically, using data that covers the dramatic increase in index investing from 2004 to 2018 we examine the monitoring behavior of index funds relative to active funds. We examine a wide variety of samples and tests, ranging from the universe of funds and firms to smaller and precisely identified subsamples. The results uniformly indicate that, relative to active funds, index funds cede power to firm management.

Our findings are consistent with the theoretical predictions in Bebchuk et al. (2017) and Edmans et al. (2018). Specifically, we find that relative to active funds, index funds are significantly more likely to side with firm management on contentious corporate governance votes. Low-fee index funds are even more likely to vote with firm management, which indicates that the low-cost structure of index funds directly affects their capacity to monitor.
Index funds are also less likely to engage with firm management either publicly or privately. In sum, our findings all point to the same conclusion: The rise of index investing is shifting power from investors to corporate managers. Bebchuk and Hirst (2019b) point out that in 2018, the three largest index fund families – Blackrock, Vanguard and State Street – cast an average of 25% of all votes in S&P 500 firms. That fraction is still increasing. This trend, along with index funds’ tendency to cede power to firm management, may be systematically weakening corporate governance.

The appropriate regulatory response is a complex question. For example, some have argued for special non-voting shares issued to index investors, or for index funds to voluntarily commit not to vote their shares. These solutions address the issue at a cost of disproportionately empowering minority shareholders. On the other end of the spectrum, Edelman, Jiang, and Thomas (2019) evaluate tenure voting systems, in which voting rights increase with the length of time that the investor holds their shares. Since index funds exit much less frequently than active funds do, such a system would tilt voting power even more toward index funds. A third option in which index funds pay a ‘monitoring fee’ as a function of their AUM with the proceeds funding an independent monitoring body might reduce the extent of the problem. Further research into index funds’ voting and monitoring behavior will be vital.
References


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Electronic copy available at: https://ssrn.com/abstract=3259433


Wei, W., & Young, A. (2019). Selection Bias or Treatment Effect? A Re-Examination of Russell 1000/2000 Index Reconstitution. *Conditionally Accepted, the Critical Finance Review*.

Figure 1. Yearly Assets Under Management for Index Funds
The figure plots the total assets under management (AUM) for passively managed index funds in the CRSP Mutual Fund database, by year, in total dollars (solid line) and as a fraction of AUM (dashed line) across all funds.
Figure 2. Index Assignment Post-2006

The figure plots assignments to the Russell 1000 and 2000 indexes in June of 2007 (vertical axis) against our proxy for Russell’s proprietary market capitalization rankings (horizontal axis). In 2007, the first year of the banding regime, stocks near the threshold all stayed in their previous years’ index, breaking the discontinuity in index assignment at rank 1000. Close to the estimated upper and lower bands (dashed lines), however, there are clear discontinuities in index switching.
Figure 3. Selection of Cohort Samples
The figure plots the sample for the 2007 cohort consisting of all Russell stocks that lay within a +/-100 rank window of the upper and lower bands, and are potential switchers, i.e. were in the Russell 2000 in 2006 for those near the upper band or were in the Russell 1000 in 2006 for those near the lower band.
Figure 4. Index Switching and Index Fund Ownership
The figure plots the evolution of index fund ownership in event time relative to index assignment. The left figure displays average ownership by Russell 2000 index funds, in event time, for firms near the lower band that were in the Russell 1000 prior to index assignment. The right figure displays average ownership by Russell 1000 index funds, in event time, for firms near the upper band that were in the Russell 2000 prior to index assignment. The vertical bars represent 95% confidence intervals.
Table I
Summary Statistics of Funds
The table presents summary statistics for all investment funds in our sample from 2004 to 2018. AUM is assets under management, in millions of USD, Expense Ratio is from the CRSP mutual fund database, and # Stocks Held is the number of stocks held in each fund on a given date.

<table>
<thead>
<tr>
<th></th>
<th>Mean</th>
<th>Std. Dev.</th>
<th>10th Pctile</th>
<th>Median</th>
<th>90th Pctile</th>
<th>Fund-Years</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Index Funds (N=1,100)</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>AUM ($M)</td>
<td>3,335</td>
<td>16,769</td>
<td>31</td>
<td>344</td>
<td>4,924</td>
<td>5,698</td>
</tr>
<tr>
<td>Expense Ratio</td>
<td>0.47%</td>
<td>0.33%</td>
<td>0.15%</td>
<td>0.43%</td>
<td>0.74%</td>
<td></td>
</tr>
<tr>
<td># Stocks Held</td>
<td>370.6</td>
<td>593.8</td>
<td>14</td>
<td>109</td>
<td>971</td>
<td>4,763</td>
</tr>
<tr>
<td><strong>Active Funds (N=4,661)</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>AUM ($M)</td>
<td>2,246</td>
<td>7,826</td>
<td>35</td>
<td>391</td>
<td>4,391</td>
<td>25,807</td>
</tr>
<tr>
<td>Expense Ratio</td>
<td>1.16%</td>
<td>0.41%</td>
<td>0.68%</td>
<td>1.12%</td>
<td>1.72%</td>
<td></td>
</tr>
<tr>
<td># Stocks Held</td>
<td>115.9</td>
<td>228.1</td>
<td>12</td>
<td>62</td>
<td>230</td>
<td>20,940</td>
</tr>
</tbody>
</table>
Table II
Summary Statistics of Firms

The table presents summary statistics for all Russell 3000 firms in Panel A and the Russell cohort sample in Panel B. Both samples run from 2004 to 2018. Firms in the Russell cohort sample are selected on the basis of lagged index membership and proximity to the upper and lower Russell bands, in each June cohort from 2007 to 2015. Each selected firm is included for three years before and after its cohort year. \textit{Market Cap} is market capitalization, in millions of USD, calculated from CRSP data. \textit{IndexOwn}^{R2000} (\textit{IndexOwn}^{R1000}) is ownership by index funds benchmarked to the Russell 2000 (1000). \textit{IndexOwn} is all ownership by passive index funds. \textit{ActiveOwn} is all ownership by active funds. \textit{TotalFundOwn} is the sum of passive and active ownership. \textit{E−Index} is the entrenchment index and ranges from 0 to 6.

Panel A: All Sample Firms

<table>
<thead>
<tr>
<th></th>
<th>Mean</th>
<th>Std. Dev.</th>
<th>10th Pctile</th>
<th>Median</th>
<th>90th Pctile</th>
<th>Firm-Years</th>
</tr>
</thead>
<tbody>
<tr>
<td>Market Cap ($M)</td>
<td>6,493</td>
<td>22,732</td>
<td>236</td>
<td>1,233</td>
<td>12,660</td>
<td>26,919</td>
</tr>
<tr>
<td>\textit{IndexOwn}^{R2000}</td>
<td>1.13%</td>
<td>1.00%</td>
<td>0.00%</td>
<td>1.34%</td>
<td>2.37%</td>
<td>26,919</td>
</tr>
<tr>
<td>\textit{IndexOwn}^{R1000}</td>
<td>0.09%</td>
<td>0.13%</td>
<td>0.00%</td>
<td>0.00%</td>
<td>0.28%</td>
<td>26,919</td>
</tr>
<tr>
<td>\textit{IndexOwn}</td>
<td>4.00%</td>
<td>2.56%</td>
<td>1.23%</td>
<td>3.49%</td>
<td>7.50%</td>
<td>26,919</td>
</tr>
<tr>
<td>\textit{ActiveOwn}</td>
<td>5.03%</td>
<td>4.23%</td>
<td>0.60%</td>
<td>4.06%</td>
<td>10.77%</td>
<td>26,919</td>
</tr>
<tr>
<td>\textit{TotalFundOwn}</td>
<td>9.03%</td>
<td>5.38%</td>
<td>2.94%</td>
<td>8.32%</td>
<td>16.07%</td>
<td>26,919</td>
</tr>
<tr>
<td>\textit{E-Index}</td>
<td>3.1</td>
<td>1.3</td>
<td>1</td>
<td>3</td>
<td>5</td>
<td>13,468</td>
</tr>
</tbody>
</table>

Panel B: Russell Cohort Firms

<table>
<thead>
<tr>
<th></th>
<th>Mean</th>
<th>Std. Dev.</th>
<th>10th Pctile</th>
<th>Median</th>
<th>90th Pctile</th>
<th>Firm-Years</th>
</tr>
</thead>
<tbody>
<tr>
<td>Market Cap ($M)</td>
<td>2,456</td>
<td>920</td>
<td>1,354</td>
<td>2,394</td>
<td>3,815</td>
<td>4,392</td>
</tr>
<tr>
<td>\textit{IndexOwn}^{R2000}</td>
<td>0.93%</td>
<td>1.00%</td>
<td>0.00%</td>
<td>0.63%</td>
<td>2.29%</td>
<td>4,392</td>
</tr>
<tr>
<td>\textit{IndexOwn}^{R1000}</td>
<td>0.09%</td>
<td>0.12%</td>
<td>0.00%</td>
<td>0.00%</td>
<td>0.27%</td>
<td>4,392</td>
</tr>
<tr>
<td>\textit{IndexOwn}</td>
<td>3.86%</td>
<td>2.60%</td>
<td>0.46%</td>
<td>3.72%</td>
<td>7.26%</td>
<td>4,392</td>
</tr>
<tr>
<td>\textit{ActiveOwn}</td>
<td>5.70%</td>
<td>4.71%</td>
<td>0.39%</td>
<td>4.78%</td>
<td>11.66%</td>
<td>4,392</td>
</tr>
<tr>
<td>\textit{TotalFundOwn}</td>
<td>9.56%</td>
<td>5.93%</td>
<td>1.58%</td>
<td>9.25%</td>
<td>16.70%</td>
<td>4,392</td>
</tr>
<tr>
<td>\textit{E-Index}</td>
<td>3.2</td>
<td>1.2</td>
<td>2</td>
<td>3</td>
<td>5</td>
<td>2,036</td>
</tr>
</tbody>
</table>
Table III
Summary Statistics of Fund Voting

The table summarizes the ISS voting data and presents comparisons of fund voting between active and index funds. The table shows the fraction of each type of fund that voted Yes, No, Abstain or that failed to vote ("did not vote", DNV) on each agenda item across all shareholder meetings of U.S. firms recorded by ISS from 2004-2018. N is the number of individual fund-vote observations.

<table>
<thead>
<tr>
<th>Management Recommend</th>
<th>ISS Recommend</th>
<th>Index funds</th>
<th>Active Funds</th>
<th>Difference</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td>Yes</td>
<td>No</td>
<td>Abstain</td>
</tr>
<tr>
<td>All</td>
<td></td>
<td>90.4%</td>
<td>6.2%</td>
<td>3.2%</td>
</tr>
<tr>
<td>Consensus</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Yes</td>
<td>Yes</td>
<td>95.6%</td>
<td>2.8%</td>
<td>1.4%</td>
</tr>
<tr>
<td>No</td>
<td>No</td>
<td>4.2%</td>
<td>84.6%</td>
<td>8.8%</td>
</tr>
<tr>
<td>Contentious</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Yes</td>
<td>No</td>
<td>54.3%</td>
<td>19.0%</td>
<td>24.9%</td>
</tr>
<tr>
<td>No</td>
<td>Yes</td>
<td>41.5%</td>
<td>53.5%</td>
<td>4.9%</td>
</tr>
</tbody>
</table>

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Table IV
Fund Ownership and Index Switching

The table presents differences-in-differences estimates of the effects of Russell index switching on investment fund ownership expressed as a percentage (1=1%) of stocks’ market capitalization. The sample consists of stocks that were “potential switchers” within a +/- 100-rank window of the yearly Russell upper and lower bands from 2007 to 2015, three years before and after index assignment for each firm in each cohort. \( R_{1000} \rightarrow R_{2000} \) equals 1 if a firm switches from the Russell 1000 to the Russell 2000, whereas \( R_{2000} \rightarrow R_{1000} \) equals 1 if a firm switches from the Russell 2000 to the Russell 1000. \( PostAssignment_t \) is an indicator variable that equals 1 after index re-balancing. The dependent variables are ownership by: (1) Russell 2000 index funds, (2) Russell 1000 index funds, (3) Index funds, (4) Active funds and (5) All mutual funds. Robust standard errors clustered by firm are shown below the estimates in parentheses. *, **, *** indicates statistical significance at the 10%, 5%, and 1% levels, respectively.

<table>
<thead>
<tr>
<th></th>
<th>( IndexOwn_{jt}^{R_{2000}} )</th>
<th>( IndexOwn_{jt}^{R_{1000}} )</th>
<th>( IndexOwn_{jt} )</th>
<th>( ActiveOwn_{jt} )</th>
<th>( TotalFundOwn_{jt} )</th>
</tr>
</thead>
<tbody>
<tr>
<td>( R_{1000} \rightarrow R_{2000} _j \times )</td>
<td>1.45***</td>
<td>-0.18***</td>
<td>1.03***</td>
<td>-0.06</td>
<td>0.97*</td>
</tr>
<tr>
<td>( PostAssignment_t )</td>
<td>(0.10)</td>
<td>(0.01)</td>
<td>(0.24)</td>
<td>(0.36)</td>
<td>(0.48)</td>
</tr>
<tr>
<td>( R_{2000} \rightarrow R_{1000} _j \times )</td>
<td>-1.34***</td>
<td>0.17***</td>
<td>-0.86***</td>
<td>-0.06</td>
<td>-0.93**</td>
</tr>
<tr>
<td>( PostAssignment_t )</td>
<td>(0.08)</td>
<td>(0.02)</td>
<td>(0.14)</td>
<td>(0.27)</td>
<td>(0.34)</td>
</tr>
<tr>
<td>Observations</td>
<td>4,392</td>
<td>4,392</td>
<td>4,392</td>
<td>4,392</td>
<td>4,392</td>
</tr>
<tr>
<td>Adjusted ( R^2 )</td>
<td>0.468</td>
<td>0.474</td>
<td>0.674</td>
<td>0.569</td>
<td>0.582</td>
</tr>
<tr>
<td>Year FE</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Firm ( \times ) Cohort FE</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
</tbody>
</table>
Table V
Fund Voting

The table presents OLS and Heckman corrected panel regression estimates examining fund voting for index funds versus active funds. Columns 1-4 show estimates for all firms in the sample. Columns 5-8 show estimates for firms that were potential switchers near the yearly Russell bands from 2007-2015. Columns 1-6 show OLS estimates, while columns 7 and 8 show estimates from a Heckman model. The dependent variable, $VotedWithMgmt$, equals 1 if a fund votes with management’s recommendation and 0 otherwise. $IndexFund$ equals 1 if the fund is an index fund, and 0 if the fund is an active fund. $ExpenseRatio$ is the fund’s total expense ratio in that year expressed in percentage points (so 25 basis points = 0.25), de-meaned for ease of interpretation. The sample consists of votes on contentious items (i.e. items on which ISS and firm management were opposed). Robust standard errors clustered by fund are shown below the estimates in parentheses. *, **, *** indicates statistical significance at the 10%, 5%, and 1% levels, respectively.

<table>
<thead>
<tr>
<th>Dependent Variable = $VotedWithMgmt$</th>
<th>(1)</th>
<th>(2)</th>
<th>(3)</th>
<th>(4)</th>
<th>(5)</th>
<th>(6)</th>
<th>(7)</th>
<th>(8)</th>
</tr>
</thead>
<tbody>
<tr>
<td>$IndexFund_i$</td>
<td>0.125***</td>
<td>0.126***</td>
<td>0.124***</td>
<td>0.126***</td>
<td>0.150***</td>
<td>0.150***</td>
<td>0.084***</td>
<td>0.117**</td>
</tr>
<tr>
<td></td>
<td>(0.025)</td>
<td>(0.024)</td>
<td>(0.024)</td>
<td>(0.024)</td>
<td>(0.030)</td>
<td>(0.030)</td>
<td>(0.032)</td>
<td>(0.059)</td>
</tr>
<tr>
<td>$ExpenseRatio_{it} \times IndexFund_i$</td>
<td>-0.238***</td>
<td>-0.238***</td>
<td>-0.209**</td>
<td>-0.228*</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.073)</td>
<td>(0.073)</td>
<td>(0.085)</td>
<td>(0.137)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>$ExpenseRatio_{it} \times$</td>
<td>0.021</td>
<td>0.017</td>
<td>0.071</td>
<td>0.065</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>$ActiveFund_i$</td>
<td>(0.046)</td>
<td>(0.045)</td>
<td>(0.060)</td>
<td>(0.064)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>$InverseMillsRatio_{ijt}$</td>
<td>-0.114***</td>
<td>0.0168</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.040)</td>
<td>(0.103)</td>
<td></td>
<td></td>
<td></td>
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</table>

<table>
<thead>
<tr>
<th>Model</th>
<th>OLS</th>
<th>OLS</th>
<th>OLS</th>
<th>OLS</th>
<th>OLS</th>
<th>OLS</th>
<th>Heckman</th>
<th>Heckman</th>
</tr>
</thead>
<tbody>
<tr>
<td>Sample Firms</td>
<td>All</td>
<td>All</td>
<td>All</td>
<td>All</td>
<td>Russell</td>
<td>Russell</td>
<td>Russell</td>
<td>Russell</td>
</tr>
<tr>
<td>Adjusted $R^2$</td>
<td>0.074</td>
<td>0.083</td>
<td>0.118</td>
<td>0.127</td>
<td>0.076</td>
<td>0.084</td>
<td>0.076</td>
<td>0.084</td>
</tr>
<tr>
<td>Firm FE</td>
<td>Yes</td>
<td>Yes</td>
<td>No</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Year FE</td>
<td>Yes</td>
<td>Yes</td>
<td>No</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Firm × Year FE</td>
<td>No</td>
<td>No</td>
<td>Yes</td>
<td>Yes</td>
<td>No</td>
<td>No</td>
<td>No</td>
<td>No</td>
</tr>
</tbody>
</table>
Table VI
Fund Voting – Split on Vote Type

The table presents OLS panel regression estimates examining fund voting for index funds versus active funds broken out by vote type. Vote type is Board of Directors in model (1), Compensation in model (2), Disclosure in model (3), and Entrenchment in model (4). The dependent variable, VotedWithMgmt, equals 1 if a fund votes with management’s recommendation and 0 otherwise. IndexFund equals 1 if the fund is an index fund, and 0 if the fund is an active fund. The sample contains the full sample of firms for votes on contentious items only (i.e. items on which ISS and firm management were opposed). Robust standard errors clustered by fund are shown below the estimates in parentheses. *, **, *** indicates statistical significance at the 10%, 5%, and 1% levels, respectively.

<table>
<thead>
<tr>
<th>Vote Type</th>
<th>Dependent Variable = VotedWithMgmt</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Board of Directors</td>
</tr>
<tr>
<td>IndexFund</td>
<td>0.132*** (0.029)</td>
</tr>
<tr>
<td>Observations</td>
<td>1,173,740</td>
</tr>
<tr>
<td>Adjusted $R^2$</td>
<td>0.086</td>
</tr>
<tr>
<td>Firm FE</td>
<td>Yes</td>
</tr>
<tr>
<td>Year FE</td>
<td>Yes</td>
</tr>
</tbody>
</table>
Table VII  
Fund Voting on Proposals by Management versus Shareholders

The table presents OLS panel regression estimates examining fund voting for index funds versus active funds splitting agenda items into items proposed by firm management (models (1) through (3)) and items proposed by shareholders (models (4) through (6)). The dependent variable is either \( VotedYes \) in models (1) and (4), \( VotedNo \) in models (2) and (5), or \( Abstained \) in models (3) and (6). \( VotedYes \) (\( VotedNo \)) equals 1 if a fund votes Yes (No) and 0 otherwise. \( Abstained \) equals 1 if a fund abstains on a vote and 0 otherwise. \( IndexFund \) equals 1 if the fund is an index fund, and 0 if the fund is an active fund. The sample contains the full sample of firms for votes on contentious items only (i.e. items on which ISS and firm management were opposed). Robust standard errors clustered by fund are shown below the estimates in parentheses. *, **, *** indicates statistical significance at the 10%, 5%, and 1% levels, respectively.

<table>
<thead>
<tr>
<th></th>
<th>Management Proposals</th>
<th>Shareholder Proposals</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>( VotedYes )</td>
<td>( VotedNo )</td>
</tr>
<tr>
<td>( IndexFund_i )</td>
<td>0.144***</td>
<td>-0.050***</td>
</tr>
<tr>
<td></td>
<td>(0.031)</td>
<td>(0.011)</td>
</tr>
<tr>
<td>Observations</td>
<td>1,408,736</td>
<td>1,408,736</td>
</tr>
<tr>
<td>Adjusted ( R^2 )</td>
<td>0.079</td>
<td>0.232</td>
</tr>
<tr>
<td>Firm FE</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Year FE</td>
<td>Yes</td>
<td>Yes</td>
</tr>
</tbody>
</table>
Table VIII
Index Switching and Changes in Agenda Items

The table presents differences-in-differences estimates of the effects of Russell index switching on the number and type of agenda items at shareholder meetings. The sample consists of stocks that were “potential switchers” within a +/- 100-rank window of the yearly Russell upper and lower bands from 2007 to 2015, three years before and after index assignment for each firm in each cohort. $R_{1000} \rightarrow R_{2000}$ equals 1 if a firm switches from the Russell 1000 to the Russell 2000, whereas $R_{2000} \rightarrow R_{1000}$ equals 1 if a firm switches from the Russell 2000 to the Russell 1000. $PostAssignment_t$ is an indicator variable that equals 1 after index re-balancing. $NumItems$ is the number of agenda items voted on in a given year. $NumShrProp$ and $NumMgmtProp$ is the number of items tabled by shareholders and management, respectively. $FracISSAgainst$ and $FracISSMgmtAgainst$ are the fraction of agenda items that were opposed by ISS and firm management respectively. $FracConsensus$ is the fraction of agenda items for which ISS and management made the same recommendation. Robust standard errors clustered by firm are shown below the estimates in parentheses. *, **, *** indicates statistical significance at the 10%, 5%, and 1% levels, respectively.

<table>
<thead>
<tr>
<th></th>
<th>NumItems, $\mu$</th>
<th>NumShrProp, $\mu$</th>
<th>NumMgmtProp, $\mu$</th>
<th>FracISSAgainst, $\mu$</th>
<th>FracMgmtAgainst, $\mu$</th>
<th>FracConsensus, $\mu$</th>
</tr>
</thead>
<tbody>
<tr>
<td>$R_{1000} \rightarrow R_{2000}$, $t$</td>
<td>0.02 (0.34)</td>
<td>-0.02 (0.07)</td>
<td>0.05 (0.32)</td>
<td>-0.01 (0.02)</td>
<td>0.00 (0.01)</td>
<td>0.01 (0.02)</td>
</tr>
<tr>
<td>$R_{2000} \rightarrow R_{1000}$, $t$</td>
<td>-0.28 (0.37)</td>
<td>0.00 (0.03)</td>
<td>-0.29 (0.37)</td>
<td>-0.00 (0.01)</td>
<td>-0.00 (0.01)</td>
<td>0.00 (0.01)</td>
</tr>
<tr>
<td>Observations</td>
<td>3,726</td>
<td>3,726</td>
<td>3,726</td>
<td>3,726</td>
<td>3,726</td>
<td>3,726</td>
</tr>
<tr>
<td>Adjusted $R^2$</td>
<td>0.614</td>
<td>0.119</td>
<td>0.623</td>
<td>0.439</td>
<td>0.031</td>
<td>0.431</td>
</tr>
<tr>
<td>Firm × Cohort FE</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Year FE</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
</tbody>
</table>
Table IX
Blockholding Disclosures: Schedule 13D versus 13G
The table presents comparisons of fund families’ blockholding disclosure filings using a probit regression. The dependent variable is “Filed 13D” which is an indicator variable for whether each filing was under the activist Schedule 13D (dependent variable = 1) versus the passive Schedule 13G (dependent variable = 0). $\text{FracAUM}_{jt}$ is the fraction of fund family $j$’s assets under management (AUM) that was managed by index funds in year $t$. $\text{logAUM}_{jt}$ is the logarithm of the fund family’s total AUM. $\text{numFilings}_{jt}$ is the number of blockholding disclosures the family filed in that year. Marginal effects are displayed in square brackets. Robust standard errors clustered by fund family are shown below the estimates in parentheses. *, **, *** indicates statistical significance at the 10%, 5%, and 1% levels respectively.

<table>
<thead>
<tr>
<th>Dependent Variable = Filed 13D</th>
<th>(1)</th>
<th>(2)</th>
<th>(3)</th>
</tr>
</thead>
<tbody>
<tr>
<td>$\text{FracAUM}_{jt}$</td>
<td>-1.13**</td>
<td>-1.05**</td>
<td>-1.15**</td>
</tr>
<tr>
<td></td>
<td>(0.48)</td>
<td>(0.46)</td>
<td>(0.49)</td>
</tr>
<tr>
<td></td>
<td>[-27%]</td>
<td>[-25%]</td>
<td>[-28%]</td>
</tr>
<tr>
<td>$\text{logAUM}_{jt}$</td>
<td>-0.052</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.042)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>$\text{numFilings}_{jt}$</td>
<td></td>
<td>0.00028</td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td>(0.00032)</td>
<td></td>
</tr>
<tr>
<td>Model</td>
<td>Probit</td>
<td>Probit</td>
<td>Probit</td>
</tr>
<tr>
<td>Observations</td>
<td>920</td>
<td>920</td>
<td>921</td>
</tr>
<tr>
<td>Pseudo $R^2$</td>
<td>0.018</td>
<td>0.018</td>
<td>0.018</td>
</tr>
</tbody>
</table>

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Table X
Index Switching and Firm Governance
The table presents differences-in-differences estimates of the effects of Russell index switching on firm governance. The sample consists of stocks that were “potential switchers” within a +/- 100-rank window of the yearly Russell upper and lower bands from 2007 to 2015, three years before and after index assignment for each firm in each cohort. $R_{1000} \rightarrow R_{2000}$ equals 1 if a firm switches from the Russell 1000 to the Russell 2000, whereas $R_{2000} \rightarrow R_{1000}$ equals 1 if a firm switches from the Russell 2000 to the Russell 1000. $PostAssignment_t$ is an indicator variable that equals 1 after index re-balancing. $E - \text{Index}$ is the entrenchment index of Bebchuk et al. (2008). $BoardIndep$ is the fraction of directors that are independent. $RemovePPill$ equals 1 if a poison pill provision was removed or allowed to lapse. $RemoveRestrict$ equals 1 if a restriction on shareholders’ ability to call a special meeting was removed or allowed to lapse. $DualClass$ equals 1 if the firm has a dual-class share structure. Robust standard errors clustered by firm are shown below the estimates in parentheses. *, **, *** indicates statistical significance at the 10%, 5%, and 1% levels, respectively.

<table>
<thead>
<tr>
<th></th>
<th>E-Index$_{jt}$</th>
<th>BoardIndep$_{jt}$</th>
<th>RemovePPill$_{jt}$</th>
<th>RemoveRestrict$_{jt}$</th>
<th>DualClass$_{jt}$</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
<td>(2)</td>
<td>(3)</td>
<td>(4)</td>
<td>(5)</td>
</tr>
<tr>
<td>$R_{1000} \rightarrow R_{2000}$$_{jt} \times PostAssignment_t$</td>
<td>-0.14*</td>
<td>-0.03***</td>
<td>-0.01</td>
<td>-0.02</td>
<td>0.00</td>
</tr>
<tr>
<td></td>
<td>(0.08)</td>
<td>(0.01)</td>
<td>(0.02)</td>
<td>(0.01)</td>
<td>(0.01)</td>
</tr>
<tr>
<td>$R_{2000} \rightarrow R_{1000}$$_{jt} \times PostAssignment_t$</td>
<td>0.06</td>
<td>0.00</td>
<td>0.04***</td>
<td>-0.02</td>
<td>0.00</td>
</tr>
<tr>
<td></td>
<td>(0.07)</td>
<td>(0.01)</td>
<td>(0.01)</td>
<td>(0.01)</td>
<td>(0.01)</td>
</tr>
<tr>
<td>Observations</td>
<td>2,469</td>
<td>2,652</td>
<td>3,876</td>
<td>3,876</td>
<td>3,876</td>
</tr>
<tr>
<td>Adjusted $R^2$</td>
<td>0.86</td>
<td>0.71</td>
<td>-0.01</td>
<td>0.16</td>
<td>0.78</td>
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<tr>
<td>Firm $\times$ Cohort FE</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
<tr>
<td>Year FE</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
<td>Yes</td>
</tr>
</tbody>
</table>

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